RESEARCH ARTICLE

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- ² Performance comparison of first-order
- conditional estimation with interaction
- and Bayesian estimation methods for
- s estimating the population parameters
- and its distribution from data sets with
 a low number of subjects

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15 Abstract

Background: Exploratory preclinical, as well as clinical trials, may involve a small number of patients, making it
 difficult to calculate and analyze the pharmacokinetic (PK) parameters, especially if the PK parameters show very
 high inter-individual variability (IIV). In this study, the performance of a classical first-order conditional estimation
 with interaction (FOCE-I) and expectation maximization (EM)-based Markov chain Monte Carlo Bayesian (BAYES)
 estimation methods were compared for estimating the population parameters and its distribution from data sets
 having a low number of subjects.

Methods: In this study, 100 data sets were simulated with eight sampling points for each subject and with six 22 different levels of IIV (5%, 10%, 20%, 30%, 50%, and 80%) in their PK parameter distribution. A stochastic simulation 23 and estimation (SSE) study was performed to simultaneously simulate data sets and estimate the parameters using 24 four different methods: FOCE-I only, BAYES(C) (FOCE-I and BAYES composite method), BAYES(F) (BAYES with all true 25 initial parameters and fixed ω^2), and BAYES only. Relative root mean squared error (rRMSE) and relative estimation 26 error (REE) were used to analyze the differences between true and estimated values. A case study was performed 27 with a clinical data of theophylline available in NONMEM distribution media. NONMEM software assisted by Pirana, 28 PsN, and Xpose was used to estimate population PK parameters, and R program was used to analyze and plot the 29 results.

Results: The rRMSE and REE values of all parameter (fixed effect and random effect) estimates showed that all four
 methods performed equally at the lower IIV levels, while the FOCE-I method performed better than other EM-based
 methods at higher IIV levels (greater than 30%). In general, estimates of random-effect parameters showed significant bias
 and imprecision, irrespective of the estimation method used and the level of IIV. Similar performance of the estimation

34 methods was observed with theophylline dataset. (Continued on next page)

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Conclusions: The classical FOCE-I method appeared to estimate the PK parameters more reliably than the BAYES method
 when using a simple model and data containing only a few subjects. EM-based estimation methods can be considered
 for adapting to the specific needs of a modeling project at later steps of modeling.

Keywords: Estimation methods, Few subjects, First-order conditional estimation with interaction, Markov chain Monte
 Carlo Bayesian, NONMEM,

40 Background

Exploratory preclinical (as well as clinical) trials may in-41 42 volve a low number of subjects (around 6 subjects). This is because in the early stages of drug development, statis-43 tical approaches are difficult to apply, potentially leading 44 to bias when predicting population mean and distribution 45 of parameters and/or all sources of variability. In addition, 46 different aspects of the study design are not considered 47 when calculating the number of subjects. As a result, it 48 can be difficult to calculate and analyze the pharmacoki-49 netic (PK) parameters, especially if the PK parameters 50 51 show very high inter-individual variability (IIV).

Population analysis is a set of statistical techniques that 52 can be used to study the average response (clinically mea-53 sured event of any biomarker) in a population, as well as 54 the IIVs in responses arising from different sources [1]. 55 NONMEM is the gold standard software for population 56 analysis that allows for mixed-effect modeling of PK/phar-57 macodynamic data while accounting for both unexplained 58 inter-subject, inter-occasion, and residual variability (ran-59 60 dom effects), as well as measured concomitant effects (fixed effects). It can also be useful for analyzing data 61 62 obtained from a low number of subjects involved in a study [2]. A list of estimation methods is available in NONMEM, 63 including classical estimation methods [first-order 64 conditional estimation with interaction (FOCE) and 65 second-order approximation (LAPLACE)] and max-66 67 imum likelihood expectation maximization (EM)-based estimation methods [iterative two-stage (ITS), import-68 ant sampling EM (IMP), important sampling EM 69 assisted by mode a posterior (IMPMAP), stochastic 70 71 approximation expectation maximization (SAEM), and 72 Markov chain Monte Carlo Bayesian (BAYES)]. Therefore, it is important to understand the performance of 73 different approach-based methods for handling data 74 with a low number of subjects. 75

Classical estimation methods like FOCE-I, including 76 77 FO, FOCE and Laplace, approximate the likelihood by taking Laplace transformation and Taylor linearization 78 [3]. These methods are known to perform well when 79 models structure are simple and low in dimension. Here, 80 the model with higher number of random-effect parame-81 82 ters (IIVs) are referred as of high dimensions. Furthermore, the classical estimation methods known to 83 provide highly reproducible values, and short run-times 84

for simple PK models [4]. However, these linearization 85 methods fail to converge and estimate parameters pre- 86 cisely with significant bias with increase in model com- 87 plexity. The EM based methods calculate the exact 88 likelihood (with approximation) by sampling and sum-89 ming through the probability density function space, 90 which is theoretically expected to approach the true like-91 lihood as the sampling reaches infinity. It is due to this 92 sampling step EM based methods have longer run-time 93 compared to the classical methods for simple PK models 94 [5]. In case of complex PK/PD problems, EM based 95 methods are faster than FOCE-I due to their efficient 96 maximization step [4]. 97

Some previous studies have compared available esti-98 mation methods with different objectives, identifying 99 various desirable traits of estimation methods. The most 100 desirable property of a given estimation method is its 101 precision and accuracy as they are the basis of the reli-102 ability of the obtained estimates. Other expected features 103 of the estimation methods are low sensitivity to priors 104 and short runtime. However, no previous study has com-105 pared estimation methods for estimating population PK 106 parameters from a small number of subjects. Therefore, 107 the objective of this study was to compare precision and 108 accuracy of estimation methods for estimating popula-109 tion mean and distribution of PK parameters from a 110 small number of subjects and explore options to 111 minimize bias with a classical method and a maximum 112 likelihood EM-based method. 113

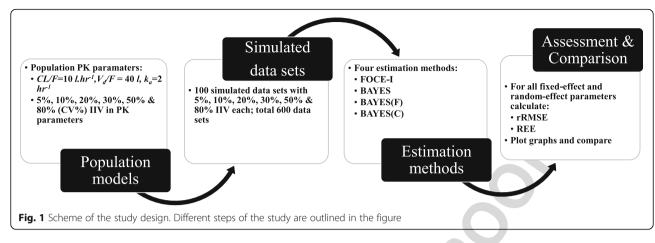
Methods

An outline of this study is provided in Fig. 1; details 115 F1 are given in the following subsections. In this study, 116 100 data sets were simulated with eight sampling 117 points for each subject and with six different levels of 118 IIV (5%, 10%, 20%, 30%, 50%, and 80%) in their PK 119 parameter distribution. The main reason for creating 120 data sets was to describe close to real situations and 121 minimize potential data set-dependent bias. 122

Stochastic simulations and estimations

A stochastic simulation and estimation (SSE) study was 124 performed using a one-compartment PK model. The 125 estimation options in the model were varied to assess 126 the performance of a classical estimation method – 127

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FOCE with the interaction option (FOCE-I), which al-128 lows for interaction between $IIV(\eta)$ and residual variabi-129 $lity(\varepsilon)$, and an EM-based estimation method – BAYES 130 estimation method in NONMEM version 7.3.0 [6] 131 assisted by Pirana (ver. 2.9.0), PsN (ver. 4.2.0), and Xpose 132 (ver. 4.4.1) [7]. For statistical analysis of the results and 133 generating different plots of the results, R (ver. 3.1.3) 134 program was used [8]. 135

136 Population model and simulated data sets

137 The population model, specifically a one-compartment open model with first-order absorption and elimination 138 rate constants, was used for simulation and estimation. 139 140 The model consisted of three systematic PK parameters 141 as fixed effects describing the absorption rate constant (K_a) , apparent volume of distribution (V_d/F) , and appar-142 ent clearance (*CL/F*), two random-effect parameters (η) 143 describing the IIV on V_d/F and CL/F [Eqs. (1, 2 and 3)], 144 and a proportional error (ϵ) model (Eq. 4): 145

$$K_a = \theta_{K_a},\tag{1}$$

$$V_d/F = \theta_{V_d/F} \cdot e^{\eta_{V_d/F}},\tag{2}$$

$$CL/F = \theta_{CL/F} \cdot e^{\eta_{CL/F}}, \tag{3}$$

$$C_{ij} = C_{pred,ij} (1 + \varepsilon_{ij}), \tag{4}$$

where C_{ij} indicates the *j*-th observations of *i*-th individual, C_{pred, ij} indicates the model-predicted C_{ij} , and ε_{ij} indicates the proportional residual error.

The following equations [Eqs. (5) and (6)] describe therate of change in drug amount in a one-compartment system:

$$\frac{dA_d}{dt} = -K_a A_d,\tag{5}$$

$$\frac{dA_c}{dt} = K_a A_d - \frac{CL/F}{V_d/F} A_c, \tag{6}$$

151 where A_d and A_c are the drug amounts in the depot and

central compartments, respectively, and *t* denotes the 152 time.

The data set used for simulation consisted of six individuals with eight sampling points within 24 h after dosing for each individual. The population mean of PK 156 parameters were assumed to be 2 L/h, 40 L, and 10 L/h 157 for $K_{av} V_d/F$, and *CL/F*, respectively and their IIV levels 158 (variance parameter ω^2) were assumed to be 5%, 10%, 159 20%, 30%, 50%, and 80% coefficient of variance (CV%) 160 (Eq. 7).

$$CV(\%) = \sqrt{e^{\omega^2 - 1} \times 100\%}.$$
 (7)

Data sets were simulated 100 times for each level of 162 IIV (total of 600 data sets) and tested to compare estimation performance in NONMEM. 164

Estimation methods

The population model was fitted to each of the simulated data sets using estimation methods with different 167 estimation options and open or fixed ω^2 values, as summarized in Table 1. 169

The FOCE-I method is a classical estimation method 170 that is applied by most users and has a short run-time 171 for estimation of population mean and distribution for 172 simple models [9, 10]. The BAYES method is a newly 173 introduced method in NONMEM and is more suitable 174 for estimation of population mean and distribution for 175 complex PK/PD models [10]. In this study, the other 176 estimation methods such as ITS, IMP, IMPMAP, and 177 SAEM were not tested because these methods were 178 expected to perform similar or below the performance 179 of BAYES as these methods are based on EM algorithms. 180 EM algorithms consist of an expectation (E) and a 181 maximization step (M), where these methods differed in 182 the way step E was performed, which involves the 183 approximation of likelihood. Additionally, the BAYES 184 method creates a large sample of probable parameters, 185

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.2		Estimation method			
.3	Methods	FOCE-I	BAYES(C)	BAYES(F)	BAYES
.4		First-order conditional estimation with interaction	FOCE-I and BAYES composite method	BAYES with ω^2 value fixed to true value	Markov chain Monte Carlo Bayesian
.5	Conditions				
.6	Initial parameters	THETAs & OMEGAs: Open true values	THETAs & OMEGAs: Open true values	THETAs: Open true values OMEGAs: Fixed true values	THETAs & OMEGAs: Open true values
.7	Estimation options	SIG = 3	For FOCE-I, SIG = 3 For BAYES, CTYPE = 3 NBURN = 4000 NITER = 10,000 SIGL = 8 NSIG = 2	CTYPE = 3 NBURN = 4000 NITER = 10,000 SIGL = 8 NSIG = 2	CTYPE = 3 NBURN = 4000 NITER = 10,000 SIGL = 8 NSIG = 2

Table 1 Estimation methods and their conditions for initial parameters and estimation options Q31.1

unlike other EM-based methods that attempt to obtain a 186 single "most likely" set of estimates. 187

In this study, true parameter values, i.e., the parameter 188 values used in the simulation step, were established as 189 initial estimates in all estimation methods. In NON-190 MEM, convergence criteria for a FOCE-I are based only 191 on the parameter estimation gradient and are tested by 192 193 default. The number of significant digits for the estimation of each parameter was set to three (SIG = 3) for the 194 195 FOCE-I method. In the BAYES estimation method, the convergence test type was set to 3 (CTYPE = 3), where 196 changes in objective function value, THETAs, OMEGAs, 197 and SIGMAs, are accessed. The number of significant 198 199 digits to which the objective function was evaluated was 200 set to 8 (SIGL = 8). In the BAYES methods, the maximum number of iterations for which to perform the 201 burn-in phase was set to 4000 (NBURN = 4000), and the 202 number of iterations for which to perform the stationary 203 distribution for BAYES analysis was set to 10,000 204 (NITER = 10,000), both of which are default values in 205 NONMEM. The former option ensured that all parame-206 ters and objective functions did not appear to move in a 207 specific direction, but appeared to instead move around 208 a stationary region, and the latter provides a large set 209 210 (10,000) of likely population parameters.

Assessment and comparison of estimation methods 211

212 The estimation methods were assessed by relative root mean squared error (rRMSE) and relative estimation 213 214 error (REE) for fixed-effect as well as random-effect parameters to calculate and visualize the magnitude of dif-215 ferences between the true value and the estimated value. 216 The rRMSE [Eq. (8)] provides a combined measure of 217 218 bias and precision.

$$rRMSE = \sqrt{\frac{\sum \frac{(P_{est} - P_{true})^2}{P_{true}}}{n^2}}$$
(8)

219 where P_{est} is the estimated parameter value, P_{est} is the

true parameter values used at the simulation step, and n 220 is the number of simulations for each set of P_{true} (*n* = 221 100). 222

REE was calculated [Eq. (9)] and plotted as box plots; 223 the plot represents the relative bias by the median of the 224 REE values and precision by distribution of REE about 225 zero.

$$REE = \sqrt{\frac{P_{est} - P_{true}}{P_{true}}}.$$
(9)

Case study

The THEO data set available in the NONMEM distribu-227 tion media was used as a case study. The estimated PK 228 parameters and IIV from the final model fitted to the 229 THEO data (called THEO model hereafter) was consid-230 ered to be the population (true) mean values for PK pa-231 rameters and IIV. SSE was performed using the THEO 232 model, where 100 data sets were simulated from the 233 model with six individuals in each data sets and four dif-234 ferent estimation methods were used, listed in Table 1, 235 to estimate the PK parameters and their IIV from the 236 100 data sets. 237

Results

The rRMSE values of the estimated parameters (fixed-ef-239 fect and random-effect) versus the level of IIV, stratified 240 based on the different PK parameters, are shown in 241 Fig. 2. The scale for each of the plots are adjusted to in-242 clude all values. The analysis of all parameter rRMSE 243 values showed that all four tested estimation methods 244 performed equally at the lower IIV levels (5–30%), while 245 the performance degraded with an increase in IIV. The 246 FOCE-I method performed better than the other three 247 EM-based estimation methods; this was more apparent 248 at higher IIV levels (above 30%) for both fixed-effect and 249 random-effect parameters. Performance of both the 250

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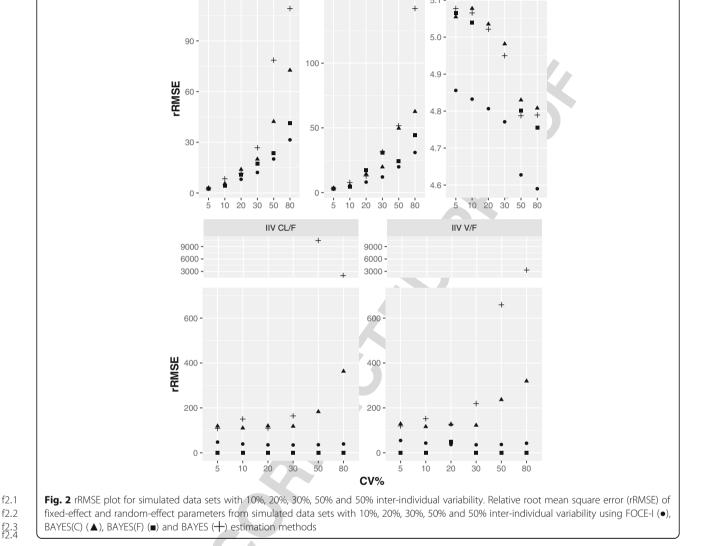
F2

BAYES(C) and BAYES methods were poor at an IIV 251 252 greater than 30% in terms of rRMSE. All parameter estimates at 50% and 80% IIV had exceptionally high 253 rRMSE. The BAYES(F) performance was intermediary 254 255 between FOCE-I and BAYES(C)/BAYES estimation methods in terms of rRMSE. 256

257 The REE of both fixed-effect and random-effect parameters versus the estimation methods, stratified by dif-258 ferent levels of IIV, are shown in Fig. 3. The plots were F3 259 adjusted to include ±100% REE for the purpose of clar-260 261 ity. In general, all estimation methods overestimated 262 fixed-effect parameters to some extent. At a lower level of IIV (5-10%), all estimation methods estimated fixed-263 effect parameters with negligible bias and reasonable 264 precision. However, the bias as well imprecision 265

increased with an increase in IIV variability. Overall, 266 FOCE-I estimated fixed-effect parameters with REE near 267 zero at all tested levels of IIV, while the distribution of 268 REE increased with an increase in IIV. The other 269 remaining three methods, BAYES(C), BAYES(F), and 270 BAYES, had comparatively higher REE with a wider dis-271 tribution range compared with the FOCE-I method. 272

The estimation of random-effect parameters had pro- 273 nounced bias and imprecision, irrespective of the estima-274 tion method used or the level of IIV (with the exception 275 of the BAYES(F) method, where the variance parameter 276 was fixed to the true value) as shown in Fig. 3. Both EM- 277 based methods, BAYES(C) and BAYES, performed poorly 278 with higher bias and impression. Across all tested levels of 279 IIV, BAYES(C) and BAYES methods had high bias and 280

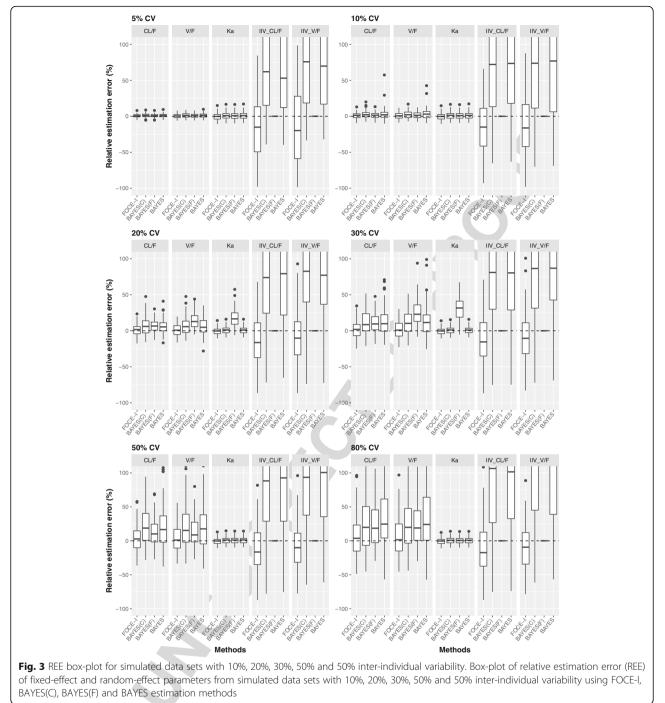


FOCE-I ▲ BAYES(C) ■ BAYES(F) + BAYES

V/F

KΔ

CL/F



f3.1 f3.2 Q4<u></u>3.3 f3.4

precision with skewed distribution of positive REE. The
FOCE-I method consistently performed better compared
with other methods with much lower and slightly negative
bias where the distribution of REE overlapped with the
zero value.

The overall stability of estimations were high with a 100% success rate of minimization and covariance step for BAYES(C), BAYES(F), and BAYES methods. For the FOCE-I method, the minimization step had a 100% success rate, but the rate of the successful covariance 290 step was 52% at 5% IIV while other estimations had a 291 successful covariance step close to 100%. 292

The THEO data set used as a case study had 132 293 observations from 12 subjects, 11 observations per individual after an oral dose of 320 mg theophylline. A one- 295 compartment PK model with first order absorption 296 described the data well and it was used as a final model. 297 The PK parameters from the THEO data set were: CL/F 298

= 2.88 l/h, V_d/F = 33.01 l and k_a = 1.46 l/h and IIV were 299 25.69%, 13.48% and 65.39%, respectively. The rRMSE 300 plots (Additional file 1) of the PK parameters from the 301 THEO model show that the performance of the four 302 estimation methods were similar for estimates of CL/F 303 and V_d/F both of which had lower IIV, below 30%. 304 Whereas, overall higher rRMSE for estimate of K_a was 305 observed, particularly from EM based methods. The esti-306 mation methods followed similar pattern of performance 307 as indicated by rRMSE for estimation of random effect 308 parameters. Similarly, REE box plots (Additional file 2) 309 for estimated PK parameters show that CL/F and V_d/F 310 estimated by all four estimation methods were very close 311 to the true values, where both of them had true IIV 312 below 30%. For the estimate of K_{a} , FOCE-I method esti-313 mated values were closes to the true value while esti-314 mated values from other three EM based method were 315 positively biased (median REE above 25%) with low pre-316 cision. Estimation of random effect parameters were 317 poor for all the estimation methods, but the FOCE-I 318 method performed relatively better in terms of bias and 319 precision. 320

321 Discussion

For an estimation method, the most desirable features 322 323 are a low bias and high precision. In this study, we used rRMSE and REE to evaluate these features. The rRMSE 324 provides a single value that indicates both bias and pre-325 cision. Moreover, rRMSE provides a way to compare 326 327 performance across parameters and models. However, the REE allows for comparison of different parameters 328 with varying magnitudes in a single plot while acknow-329 ledging bias and precision. For an estimation method to 330 be unbiased and precise, the REE should have a normal 331 332 distribution with a median of 0 and a narrow range of values. 333

The FOCE-I method performed better among the four 334 methods tested based on the overall rRMSE. This per-335 formance was supported by the REE plot, which did not 336 show any significant bias for any fixed effect parameters 337 at any given level of IIV. The median REE values for the 338 random-effect parameters were not greater than -17% at 339 any given level of IIV. A resembling result of negative 340 341 bias was observed with the FOCE-I algorithm in a similar studies comparing different estimation methods [9]. 342 343 The FOCE-I method has been shown to work sufficiently well for simple models when compared to other 344 EM based algorithms in previous studies. Furthermore, 345 when the IIV was low, the performance of classical esti-346 347 mation methods and EM based methods were very close. 348 Similar results were observed in a previous study for such simple model (1-compartment model), where the 349 performance of those estimation methods were found to 350 be nearly equal [5]. 351

On the other hand, rRMSE values for the three BAYES-352 based methods were significantly higher for both fixed-353 and random-effect parameters at higher levels of IIV. The 354 higher rRMSEs were due to the wider spread of outliers, 355 more so at higher levels of IIV. A similar trend of rRMSE 356 of estimated parameters was observed using BAYES 357 methods by Johansson et al., where the highly distorted 358 rRMSE rendered the estimated parameters meaningless 359 [9]. The performances of the BAYES-based methods were 360 poor, with high bias and low precision. Even with the 361 utilization of true values for all initial parameters, the 362 BAYES(F) method was not able to estimate parameters 363 close to the true values. Similarly, the median REE for all 364 three methods based on the BAYES-method was com-365 paratively higher for fixed-effect parameters and signifi-366 cantly higher for random-effect parameters, compared 367 with those of the classical FOCE-I method. There was also 368 a general trend of an increase in REE (positive) with an 369 increase in IIV. These observations with BAYES-based 370 methods can be attributed to the way in which the BAYES 371 method estimates the parameters i.e., by generating a large 372 set of probable population parameters and variance 373 parameters that represent the distribution according to 374 their ability to fit the data [11]. Therefore, the limited 375 number of subjects used in the study may be the reason 376 for the poor performance of the three BAYES-based 377 methods. However, a previous study showed that the 378 BAYES method can provide robust estimates of complex 379 PK/PD models with rich data and reliable priors [10]. 380

The classical estimation method, FOCE-I, and max-381 imum likelihood EM-based BAYES method differ in 382 their convergence criteria, where the former is based on 383 changes in the parameter estimation gradient and are 384 tested by default, and the latter is based on changes in 385 objective function value and parameter estimates. The 386 BAYES method can also define the convergence test 387 type, and one can choose from no test, tests accessing 388 changes in objective function, thetas and sigmas only, 389 the addition of diagonals of omegas, or the addition of 390 all omegas. For these reasons, the convergence rate was 391 not included as a factor for comparison of estimation 392 methods. However, all four methods tested at any level 393 of IIV showed a 100% convergence rate. Additionally, in 394 all estimation methods, the default or generally used 395 values were used for options in the \$ESTIMATION 396 block. It is possible to optimize the outcomes by chan-397 ging the values for different options in \$ESTIMATION 398 block [4]. However, this aspect of the estimation method 399 was not compared, as this study only explored the prac-400 tice of most users. 401

In this study, FOCE-I, the classical method, performed 402 better with lower bias and higher precision compared 403 with other BAYES-based methods. Moreover, the FOCE- 404 I method is known to have a shorter run time that any 405

other new methods [9, 12]. The work presented here 406 compares a classical estimation method, FOCE-I, and 407 BAYES method, with different options in the \$ESTIMA-408 TION block and fixed OMEGA values (BAYES(C), 409 BAYES(F), and BAYES) for population analysis of data 410 with a low number of subjects (n = 6). Moreover, the 411 models built had only one compartment, with basic PK 412 parameters and random effects on two PK parameters. 413 Therefore, it should be noted that the structure and 414 complexity of a model might vary (increase or decrease) 415 in different pharmacometric projects or within the same 416 project from the initial to final step. In contrast to our 417 study, other studies have shown that for complex models 418 with highly non-linear functions [12], highly skewed 419 count distributions [13, 14], and/or low variability or 420 very rare events [15], the classical methods exhibit 421 marked bias and impression. Additionally, the selection 422 of an estimation method for a particular modeling pro-423 ject can depend on various aspects including bias, preci-474 sion, robustness, runtime, data type, timeframe of 425 project, application of results etc. which are objective in 426 nature as well as subjective aspect such as preference for 427 particular estimation method based on knowledge and 428 previous experience. Ultimately, a pharmacometrician 429 needs to make a choice for an estimation method based 430 431 on multiple aspects.

The data sets used in this study, unlike real clinical 432 data, were simulated. IIV for all PK parameters were 433 assumed to be the same for an individual; i.e., IIV 434 was either 5%, 10%, 20%, 30%, 50%, or 80% for K_{av} 435 436 CL/F_{t} and V_{d}/F_{t} . In clinical scenarios, the CV may vary widely among the PK parameters within an indi-437 vidual. Therefore, to access the relevance of results 438 obtained from simulated data, a clinical data of theo-439 phylline involving 12 subjects, THEO data set, was 440 used as a case study. The limitation of using real data 441 is that the expected true parameters value is un-442 known. So, SSE was performed, where the parameter 443 estimates from final THEO model was considered to 444 be true parameters. And the parameter estimates 445 from different estimation methods were compared to 446 so-called true values for compare their performance. 447 Similarity in the performance of all four estimations 448 449 methods at lower IIV and better performance of FOCE-I methods at higher IIV was demonstrated by 450 451 the rRMSE and REE of the estimated parameters. This further supports the results from the simulated 452 data. Another limitation of this study is that only 453 FOCE-I and BAYES methods were tested and com-454 455 pared. To further explore the best estimation method 456 when dealing with a low number of subjects, other methods in NONMEM, such as LAPLACE, ITS, IMP, 457 IMPMAP, and SAEM should also be evaluated in 458 future studies. 459

Conclusions

The FOCE-I, a classical estimation method, yielded better 461 results in terms of bias and precision across all levels of IIV 462 in comparison to three variations of BAYES estimation 463 methods. The difference in performance between FOCE-I 464 and three BAYES estimation methods in estimating fixed-465 effect parameters were significant only at the IIV level 466 greater than 30%. The bias and imprecision of random-467 effect parameters were higher compared with fixed-effect 468 parameters, however, it was consistently lower for FOCE-I 469 method compared to those estimated using BAYES(C) and 470 BAYES methods. These results were further supported by 471 the results from the THEO data, where clinical data was 472 used to simultaneously simulate and estimate PK parame-473 ters using FOCE-I and three BAYES estimation methods. 474

In conclusion, the classical FOCE-I method estimated 475 the PK parameters more reliably than the BAYES 476 method when using a simple model and data containing 477 only a few subjects. After the base modeling step is 478 complete and/or at the pivotal modeling step, use of 479 other EM-based estimation methods can be considered 480 for adapting to specific needs of the project. 481

Additional files

Additional file 1: rRMSE plot for THEO data set. Relative root mean square error (rRMSE) of fixed-effect and random-effect parameters from THEO data set using FOCE-I (●), BAYES(C) (▲), BAYES(F) (■) and BAYES (+) estimation methods. (PDF 6 kb)

Additional file 2: REE box for THEO data set. Box-plot of relative estimation error (REE) of fixed-effect and random-effect parameters from THEO data set using FOCE-I, BAYES(C), BAYES(F) and BAYES estimation methods. (PDF 8 kb)

Abbreviations

BAYES: Markov chain Monte Carlo Bayesian; BAYES(C): First-order conditional 494 estimation with interaction and Markov chain Monte Carlo Bayesian 495 composite method: BAYES(F): Markov chain Monte Carlo Bavesian with 496 variance parameter fixed to true value; CL/F: Apparent clearance; &: Residual 497 498 variability; EM: Expectation maximisation; FOCE: First-order conditional estimation; FOCE-I: First-order conditional estimation with interaction; 499 IIV: Inter-individual variability; IMP: Important sampling; IMPMAP: Important 500 sampling assisted by mode a posterior; ITS: Iterative two-stage; 501 Ka: Absorption rate constant; n: Inter-individual variability; 502 PK: Pharmacokinetic: REE: Relative estimation error: rRMSE: Relative root mean 503 square error; SAEM: Stochastic approximation expectation maximization; 504 SSE: Stochastic simulation and estimation: Vd/F: Apparent volume of 505 distribution 506

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University. 517

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519 Availability of data and materials

- 520 Data used in this study was simulated in NONMEM using SSE tool. THEO
- 521 data used as a case study is available in the NONMEM distribution media.

522 Authors' contributions

- 523 SP and BS made substantial contributions to conception and design of the
- 524 study and performed the simulation and estimation part with assistance
- 525 from JL. SP and BS drafted the manuscript. JC, KIK, HB and NH assisted in
- 526 critical analysis and interpretation of data as well as revising the contents
- 527 critically for every draft of the manuscript. This work was performed under
- 528 direct supervision and guidance of KK and HY, they guided the team from
- 529 initial concept of the project through many iterative modifications in
- 530 methods, data analysis and interpretation, critically revising every draft of
- 531 manuscript and approving final version to be submitted. All authors agree to
- 532 be accountable for all aspects of the work in ensuing the questions related 533 to the accuracy or integrity of any part of the work are appropriately
- investigated and resolved. 534

535 Ethics approval and consent to participate

536 Not applicable.

537 Consent for publication

- 538 Not applicable.

539 Competing interests

540 The authors declare that they have no competing interests.

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