RESEARCH ARTICLE

Generative adversarial networks for imputing missing data for big data clinical research

Weinan Dong¹, Daniel Yee Tak Fong², Jin-sun Yoon³, Eric Yuk Fai Wan^{1,4*}, Laura Elizabeth Bedford¹, Eric Ho Man Tang¹ and Cindy Lo Kuen Lam¹

Abstract

Background: Missing data is a pervasive problem in clinical research. Generative adversarial imputation nets (GAIN), a novel machine learning data imputation approach, has the potential to substitute missing data accurately and efficiently but has not yet been evaluated in empirical big clinical datasets.

Objectives: This study aimed to evaluate the accuracy of GAIN in imputing missing values in large real-world clinical datasets with mixed-type variables. The computation efficiency of GAIN was also evaluated. The performance of GAIN was compared with other commonly used methods, MICE and missForest.

Methods: Two real world clinical datasets were used. The first was that of a cohort study on the long-term outcomes of patients with diabetes (50,000 complete cases), and the second was of a cohort study on the effectiveness of a risk assessment and management programme for patients with hypertension (10,000 complete cases). Missing data (missing at random) to independent variables were simulated at different missingness rates (20, 50%). The normalized root mean square error (NRMSE) between imputed values and real values for continuous variables and the proportion of falsely classified (PFC) for categorical variables were used to measure imputation accuracy. Computation time per imputation for each method was recorded. The differences in accuracy of different imputation methods were compared using ANOVA or non-parametric test.

Results: Both missForest and GAIN were more accurate than MICE. GAIN showed similar accuracy as missForest when the simulated missingness rate was 20%, but was more accurate when the simulated missingness rate was 50%. GAIN was the most accurate for the imputation of skewed continuous and imbalanced categorical variables at both missingness rates. GAIN had a much higher computation speed (32 min on PC) comparing to that of missForest (1300 min) when the sample size is 50,000.

(Continued on next page)

* Correspondence: yfwan@hku.hk

¹Department of Family Medicine and Primary Care, Faculty of Medicine,

University of Hong Kong, Hong Kong, Hong Kong SAR, China

⁴Department of Pharmacology and Pharmacy, Faculty of Medicine, University

of Hong Kong, Hong Kong, Hong Kong SAR, China

BMC

Full list of author information is available at the end of the article

appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/. The Creative Commons Public Domain Dedication waiver (http://creativecommons.org/publicidomain/zero/1.0/) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

© The Author(s), 2021 Open Access This article is licensed under a Creative Commons Attribution 4.0 International License.

which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give



Open Access

(Continued from previous page)

Conclusion: GAIN showed better accuracy as an imputation method for missing data in large real-world clinical datasets compared to MICE and missForest, and was more resistant to high missingness rate (50%). The high computation speed is an added advantage of GAIN in big clinical data research. It holds potential as an accurate and efficient method for missing data imputation in future big data clinical research.

Trial registration: ClinicalTrials.gov ID: NCT03299010; Unique Protocol ID: HKUCTR-2232

Keywords: Generative adversarial network, Missing data imputation, Machine learning, Clinical research, Big data

Background

Missing data is a pervasive problem in big data research, clinical trials and epidemiological studies [1]. There are a number of reasons that could account for missing data, such as non-response to questionnaires, study participants lost to follow up, omission of data entry, failure of equipment, or incomplete or lost records [2, 3]. Mere exclusion of cases with missing data from analysis may lead to biased inference, reduced statistical power and generalisability of results [4, 5]. According to missingness assumptions, the problem of missing data can be classified into three categories, including missing completely at random (MCAR), missing at random (MAR), and missing not at random (MNAR) [6-8]. In general, the majority of the missing data in medical research are assumed to be MAR [9]. In contrast to MCAR, where there are no systematic differences between the missing and observed values, with MAR data, there will be differences between missing and observed values but these differences can be explained by other observed data [10–12].

Multiple imputation by chained equations (MICE) is the most commonly used statistical procedure for handling missing data [5], particularly for data that are MAR [13]. MICE is widely available in many statistical software including SPSS, STATA and R. Although it is important to note that MICE may lead to biased results because, by default it uses predictive mean matching (pmm) and logistic regression (LR), which are limited in the ability to handle non-linear relationships and interactions between variables [14]. A mean to overcome non-linearity is through random forest, an ensemble machine learning algorithm of multi-classification or decision tree regression [15]. Stekhoven et al. have developed a method 'missForest' (based on random forest) to impute missing values in mixed-type datasets [16]. Subsequent studies have shown that missForest outperformed MICE in both simulated and real world datasets [17]. However, a drawback of missForest is that its long computation time, limiting its practicality in big data research.

Generative adversarial network (GAN), an unsupervised algorithm, is a popular machine learning method that has been widely applied in both data generation [18] and image processing [19]. Generative adversarial imputation nets (GAIN), which is based on GAN, was recently developed and found to outperform other methods in terms of imputation accuracy in substituting MCAR data in five open-source datasets [20]. However, the accuracy of GAIN for imputing MAR data and of mixed-type variables, both of which are common in medical research remains unclear.

The main aim of this study was to evaluate the accuracy of GAIN in imputing missing values in realworld clinical datasets with mixed-type variables. Further, this study also aimed to examine the computation efficiency of GAIN as well as compare its performance with those of MICE and missForest. It is anticipated that the results will inform other researchers on the choice of missing data imputation methods in big clinical data research.

Methods

Study setting and datasets

Two large real world clinical datasets from two longitudinal cohort studies on primary care patients with chronic diseases were used. The first dataset was that of a study on the prediction of complications and mortality among a cohort of 141,516 patients with diabetes [21]. A total of 14 (out of 21) independent baseline variables had missing data, of which 12 variables had a missingness rate of less than 20%. Overall, the proportion of missing data ranged from 0.50% (systolic blood pressure) to 48.99% (urine albumin to creatinine ratio [Urine ACR]). Urine ACR showed the highest proportion of missing data as it was not routinely collected in Hong Kong primary care prior to 2010. We selected 50,000 subjects without any missing values for these 21 variables (15 continuous predictors and six categorical predictors) at baseline and seven dependent outcome variables measuring various complications of diabetes and mortality.

The second dataset was that of a cohort study evaluating the effectiveness of a risk assessment and management programme for patients with hypertension [22]. We identified 10 independent variables, including five continuous variables and five categorical variables, for inclusion in the analyses. In the original dataset, the data missingness rate for these 10 variables ranged from 1.5 to 26%. A total of 10,000 subjects without any missing values for these 10 variables were randomly selected. The data were extracted together with the data for the two outcome variables (cardiovascular diseases [CVD] and mortality) in order to replicate the imputation analyses and strengthen the generalizability of the results from the first dataset.

For easy reference, the first dataset is referred as the 'DM-data' and the second is referred as the 'HT-data'. The description of the characteristics for these two datasets can be found in Supplementary Tables 1 & 2.

Institutional Review Board of the University of Hong Kong—the Hospital Authority Hong Kong West Cluster (reference number: UW 15–258) approved this study and usage of data. Individualized informed consent is not required. All methods on the datasets were carried out in accordance with relevant guidelines and regulations.

Missing data simulation

For both DM-data and HT-data, data 'missing at random' (MAR) was simulated at different missingness rates (20 and 50%) to create the datasets for the imputation testing [17, 23]. The missingness was introduced to independent variables following Bernoulli distributions based on linear combination of dependent variables (fullyobserved). At each missingness rate, ten different incomplete datasets were generated using different randomised linear combination parameters. We did not simulate missing values in the dependent variables, although they were incorporated in the imputation process as auxiliary variables [24].

Imputation procedures with GAIN

A number of improvements were applied to the basic GAIN construction built by Yoon et al. [20] to optimize the model. First, the random noise was substituted by the mean value of each variable so as to reach the optimal solution faster. Batch normalization with gradient descent optimizer was also used to allow a larger learning rate. Combination of the loss of continuous and categorical variables with separate weights (α and β) was used to deal with a dataset with mixed types of variables. A greedy search strategy was adopted to seek the best combination of hyper-parameters. This strategy was adopted due to the large number of hyperparameters to be tuned in the GAIN training process, including k, p_{hint} , α , β , number of iterations, number of hidden layers, number of neurons in each layer, activation functions, learning rate and optimizer. The code is available at Github (https://github.com/ dongdongdwn/GAIN-Dovey) and the optimal hyper-parameters are presented in the Supplementary

Table 3. The brief imputation procedures with GAIN are presented in Algorithm 1.

Algorithm 1 Implementation of GAIN				
Given a to be imputed dataset X, M is a same shape matrix to identify the missing values, Z is a same shape matrix presenting the mean value of each row vector in X.				
1. All the variables were standardized to 0-1, lest different scales affect variables' weights.				
2. Build a fully connected multilayer neural network generator G, taking X, M and Z as inputs. The output of G in the initial training can be written as:				
$\overline{X} = G((X \odot M + Z \odot (1 - M)), M)$				
Following training can be written as:				
$\hat{X} = G((X \odot M + \overline{X} \odot (1 - M)), M)$				
3. Build a fully connected multilayer neural network discriminator. The inputs of D were the imputed dataset by generator and a hint matrix H. H was created by hint mechanism to provide more information for D to operate better discrimination, which can be written as:				
$h \in H = B \odot M + 0.5 (1 - B)$				
where B is a random variable $B = (B_1,, B_d) \in \{0, 1\}^d$ Then D output the possibility that a value is real or fake: $D_{output} = D(\hat{X}, H)$				
4. Adversarial nets and training k samples from the dataset were randomly extracted as mini-batch taking part in the training process. Starting from training D, above D and G were alternately trained to optimize respective loss function as below: $D_{loss} = \max_{D} [M \log D(\hat{X}, H) + (1 - M) \log(1 - D(\hat{X}, H))]$				

 $G_{loss} = \min_{c} (L_G + \alpha L_{M1} + \beta L_{M2})$

 $L_G = -(1 - M) \log D(\hat{X}, H)$

where L_{M1} is the sum of mean square error (MSE) between the observed values and predicted values at the same location, but specific for continuous variables, L_{M2} is the sum of cross-entropy loss of counterpart for categorical variables, α and β are hyper-parameters.

5. Inverse standardization to original scale

MICE and missForest

Imputation by MICE and missForest were carried out by standard procedures [16, 24] with R package *mice v3.6.0* and *missForest*. The imputation model of MICE was specified as predictive mean matching (pmm) and logistic regression (LR) as default, respectively, for continuous variables and categorical variables. The iteration number was set to 10. For missForest, the number of trees was set to 20, and the number of variables randomly sampled at each split was set to $d^{\frac{1}{2}}$ (sqrt dimensionality). The maxiteration number of missForest was set to 10. The iteration numbers of MICE and missForest were determined based on preliminary experiments to ensure they could achieve the best performance (as shown in Supplementary Fig. 1).

Outcome measures and data analysis

Accuracy was measured by imputation error, defined as the difference between the imputed values and real values. It was assessed by normalized root mean square error (NRMSE) for continuous variables and proportion of falsely classified (PFC) subjects for categorical variables. NRMSE and PFC were defined as follows:

NRMSE =
$$\frac{\sqrt{\frac{1}{N}\sum_{i=1}^{N}(\widehat{x}_{i}-x_{i})^{2}}}{\frac{\sum_{i=1}^{N}x_{i}}{N}}$$
PFC = $1 - \frac{N_{correct}}{N}$

where $\hat{x_i}$ is the imputed value and x_i is the original value in continuous variables, $N_{correct}$ is the total number of correctly classified values in categorical variables.

For each simulated incomplete dataset, the imputation was repeated 100 times using each method. The mean NRMSE for each continuous variable was calculated by averaging the NRMSE obtained from the 100 imputations. The mean PFC was calculated by averaging the PFC obtained in each imputation for categorical variables. NRMSE and PFC were treated as continuous variables in the comparative analysis, and their distributions were tested by Shapiro-Wilk normality test. Correspondingly, the differences in mean NRMSE or PFC among methods were tested by oneway ANOVA or non-parametric test.

Density plots and bar plots were used to illustrate the imputation differences among methods, for representative continuous variables and categorical variables respectively. For the DM-data, systolic blood pressure (SBP), fasting glucose, hypertension history and smoking status were selected to represent normal distributed continuous variables, skewed continuous variables, balanced categorical variables and imbalanced categorical variables, respectively. Likewise, age, total cholesterol to high-density lipoprotein (TC/HDL) ratio, sex and lipid lowering drugs usage were selected as the representative variables for the HT-data.

For DM-data with 5000 to 50,000 subjects, the computation time of each method to complete an imputation process on a personal computer (PC) and high performance computing (HPC) device was recorded and plotted for comparison. The relevant machine configuration of the PC and HPC can be found in Supplementary Table 3.

Missing data simulation, MICE, missForest and comparison were operated in R 3.5.1. GAIN was developed with Python 3.5. The level of significance for all statistical tests was set as 0.05.

Results

Experiments on DM-data

Table 1 presents the imputation errors (NRMSE and PFC for continuous and categorical variables, respectively) of different imputation methods at missingness rates of 20 and 50%. Overall, GAIN and missForest were superior to MICE for both continuous and categorical variables, irrespective of the missingness rates (p < 0.001). When the missingness rate was 20%, GAIN was superior to missForest with lower imputation errors (p < 0.05) for highly skewed (skewness>4) continuous variables (e.g., creatinine, fasting glucose, urine ACR) and highly imbalanced categorical variables (proportion of minority class was close to or lower than 10%, e.g., lipid lowering drug usage, DM treatment). MissForest showed better accuracy for some normally distributed continuous variables (e.g., age, SBP, DBP) and some relatively balanced categorical variables (e.g., sex, hypertension history) (p < 0.05). GAIN and missForest showed similar accuracy for the remaining variables (p > 0.05). However, GAIN was superior to missForest for the majority of variables when the missingness rate increased to 50% (p < 0.05). No statistically significant differences were observed between GAIN and missForest for the less skewed continuous variables (e.g., age, SBP, LDL-C) and relatively balanced categorical variables (e.g., sex, hypertension history).

Experiments on HT-data

The imputation errors in the HT-data of different methods are presented in Table 2. The findings were similar to those found in the DM-data. Overall, GAIN and missForest outperformed MICE for both missingness rates (20 and 50%) irrespective of the type of variables. When the missingness rate was 20%, GAIN was superior to missForest for more skewed continuous variables (e.g., SBP, TC/HDL-C ratio, hospital admission times) and more imbalanced categorical variables (e.g., smoking, hypertensive drugs, lipid lowering drugs). If the missingness rate increased to 50%, GAIN was more accurate than missForest for the majority of the variables (p < 0.05).

To illustrate the differences of the imputation errors among methods, density plots and bar plots were used to visualize the representative variables at 50% missingness rate. Density plots, showing the distribution of the absolute difference between imputed values and real values of continuous variables, are presented in Fig. 1. The absolute differences between real values and values generated by GAIN were more close to 0 and concentrated, indicating good accuracy. MICE tended to have a broader distribution of errors and a higher density of greater errors. The differences in the patterns among different methods were more noticeable on data that were skewed (e.g. fasting glucose, TC/HDL ratio).

The bar plots illustrate the distribution of imputed values and the correct proportion in each category (Fig. 2). The imputed values of MICE and GAIN showed the same distribution as the original data, while missForest generated a higher proportion of the majority group but a lower proportion of the minority group. Meanwhile, for both balanced (i.e. sex, hypertension history) and imbalanced categorical variables (i.e. smoking, lipid lowering drugs usage), GAIN imputation resulted in a more accurate allocation to the minority group when compared to the other two methods.

Computation time

The computation time of one imputation process on the DM-data by each method using PC and HPC for different sample sizes are presented in Fig. 3. MICE was the fastest for small sample sizes (up to 30,000

thods in DM-data	
mputation errors of different method	
n errors of	
Imputation	
Table 1	

	Skewness or proportion	Missingness rate = 20%	= 20%		Missingness rate = 50%	: = 50%	
	of minority class	MICE	missForest	GAIN	MICE	missForest	GAIN
Continuous variables							
Age, years	-0.106	0.078 ± 0.002	$0.060 \pm 0.001^{a,b}$	0.069 ± 0.002 ^a	0.137 ± 0.002	0.107 ± 0.001 ^a	0.111 ± 0.002^{a}
SBP, mmHg	0.316	0.052 ± 0.001	$0.041 \pm 0.002^{a,b}$	0.048 ± 0.002^{a}	0.099 ± 0.001	0.082 ± 0.002^{a}	0.080 ± 0.002 ^a
DBP, mmHg	0.154	0.070 ± 0.002	$0.052 \pm 0.002^{a,b}$	0.056 ± 0.002 ^a	0.120 ± 0.002	0.094 ± 0.001^{a}	$0.090 \pm 0.001^{-a,c}$
LDL-C, mmol/L	0.379	0.095 ± 0.003	0.075 ± 0.003 ^{a,b}	0.089 ± 0.003 ^a	0.208 ± 0.003	0.163 ± 0.004 ^a	0.161 ± 0.004 ^a
BMI, kg/m2	0.813	0.064 ± 0.003	0.048 ± 0.002^{a}	0.048 ± 0.003 ^a	0.120 ± 0.004	0.095 ± 0.002^{a}	0.090 ± 0.005 ^{a,c}
Waist, cm	0.299	0.047 ± 0.002	0.036 ± 0.001 ^a	0.036 ± 0.001^{a}	0.088 ± 0.002	0.069 ± 0.001^{a}	0.067 ± 0.003 ^a
TC, mmol/L	0.564	0.065 ± 0.003	$0.050 \pm 0.002^{a,b}$	0.055 ± 0.003 ^a	0.140 ± 0.003	0.110 ± 0.003 ^a	0.102 ± 0.004 ^{a,c}
DM duration, years	-1.167	0.284 ± 0.007	0.206 ± 0.006^{a}	0.190 ± 0.006 ^{a,c}	0.451 ± 0.01	0.340 ± 0.006 ^a	$0.304 \pm 0.012^{a,c}$
eGFR, ml/min/1.73 m2	1.368	0.089 ± 0.010	0.057 ± 0.004 ^{a,b}	0.087 ± 0.006	0.195 ± 0.012	0.159 ± 0.012 ^a	0.146 ± 0.015 ^{a,c}
HbA1c, %	1.557	0.106 ± 0.004	0.077 ± 0.002^{a}	0.078 ± 0.004 ^a	0.177 ± 0.007	0.138 ± 0.004 ^a	0.125 ± 0.004 ^{a,c}
HDL-C, mmol/L	2.729	0.132 ± 0.016	0.111 ± 0.014 ^a	0.115 ± 0.011 ^a	0.251 ± 0.011	0.197 ± 0.014 ^a	0.184 ± 0.015 ^{a,c}
TG, mmol/L	3.932	0.287 ± 0.027	0.251 ± 0.022 ^a	0.266 ± 0.027	0.610 ± 0.027	0.486 ± 0.023 ^a	0.444 ± 0.026 ^{a,c}
Creatinine, µmol/L	4.128	0.093 ± 0.011	0.089 ± 0.016	$0.068 \pm 0.015^{a,c}$	0.218 ± 0.013	0.177 ± 0.019 ^a	0.169 ± 0.015 ^{a,c}
Fasting glucose, mmol/L	4.681	0.178 ± 0.043	0.121 ± 0.008 ^a	$0.118 \pm 0.007^{a,c}$	0.277 ± 0.024	0.214 ± 0.011^{-3}	$0.195 \pm 0.010^{-a,c}$
Urine ACR, mg/mmol	11.450	2.509 ± 0.441	1.728 ± 0.307 ^a	1.554 ± 0.266 ^{a,c}	3.843 ± 0.405	2.987 ± 0.240 ^a	2.690 ± 0.258 ^{a,c}
Categorical variables							
Lipid drug usage	8.50%	0.162 ± 0.013	0.093 ± 0.010^{a}	$0.083 \pm 0.009^{-a,c}$	0.159 ± 0.006	0.090 ± 0.004 ^a	$0.079 \pm 0.005^{a,c}$
Smoker	10.57%	0.176 ± 0.014	0.113 ± 0.010^{a}	$0.094 \pm 0.009^{a,c}$	0.182 ± 0.013	0.122 ± 0.007 ^a	0.097 ± 0.008 ^{a,c}
DM treatment	10.50%	0.179 ± 0.013	0.115 ± 0.009^{a}	0.095 ± 0.009 ^{a,c}	0.187 ± 0.011	0.120 ± 0.006 ^a	0.096 ± 0.003 ^{a,c}
Hypertension drug usage	29.68%	0.318 ± 0.020	0.256 ± 0.015^{a}	0.267 ± 0.016 ^a	0.345 ± 0.01	0.281 ± 0.011 ^a	0.274 ± 0.013 ^{a,c}
Sex	45.93%	0.205 ± 0.020	$0.126 \pm 0.009^{a,b}$	0.235 ± 0.027	0.353 ± 0.011	0.276 ± 0.01 ^a	0.287 ± 0.014 ^a
Hypertension history	47.190%	0.122 ± 0.011	0.077 ± 0.008 ^{a,b}	0.129 ± 0.040	0.255 ± 0.012	0.201 ± 0.019 ^a	0.215 ± 0.017 ^a
Notes							

SBP Systolic Blood Pressure, DBP Diastolic Blood Pressure, LDL-C Low Density Lipoprotein-Cholesterol, BMI Body Mass Index, TC Total Cholesterol, eGFR Estimated Glomerular Filtration, HbA1c Hemoglobin A1c, HDL-C High Density Lipoprotein-Cholesterol, TG Triglyceride, Urine ACR Urine Albumin to Creatinine Ratio Since NRMSE and PFC both followed normal distribution (Shapiro-Wilk normality test p value > 0.05), imputation errors of different methods were compared using one-way ANOVA. If p < 0.05, paired methods were

compared using independent sample t-test ^aThe mean imputation error is significantly lower than that of MICE (p < 0.05) ^bThe mean imputation error is significantly lower than that of GAIN (p < 0.05) ^cThe mean imputation error is significantly lower than that of missForest (p < 0.05)

Table 2 Imputation errors of different methods in HT-data

	Skewness or proportion of minority class	MICE	missForest	GAIN
Missingness rate = 20%				
Continuous variables				
Age, years	-0.018	0.063 ± 0.002	0.049 ± 0.001 ^{a,b}	0.057 \pm 0.004 $^{\rm a}$
SBP	0.492	0.075 ± 0.001	0.058 ± 0.000 ^a	0.048 ± 0.000 ^{a,c}
Charlson index	0.146	0.154 ± 0.002	0.121 ± 0.001 ^{a,b}	0.144 \pm 0.003 $^{\rm a}$
TC/HDL-C ratio	3.139	0.175 ± 0.003	0.137 ± 0.001 ^a	0.115 ± 0.001 ^{a,c}
Hospital admission times	7.037	2.379 ± 0.069	1.885 ± 0.042 ^a	1.752 ± 0.141 ^{a,c}
Categorical variables				
Smoking	7.45%	0.133 ± 0.007	0.123 ± 0.003 ^a	$0.098 \pm 0.010^{-a,c}$
Hypertensive drugs	8.10%	0.149 ± 0.006	0.126 ± 0.003 ^a	0.098 ± 0.002 ^{a,d}
Lipid Lowering drugs	9.99%	0.173 ± 0.007	0.159 ± 0.003 ^a	0.129 ± 0.006 ^{a,c}
Overweight	37.89%	0.433 ± 0.01	0.400 ± 0.005 $^{\rm a}$	0.359 ± 0.003 ^{a,}
Sex	41.21%	0.448 ± 0.019	0.412 ± 0.004 ^a	0.405 ± 0.022 ^a
Missingness rate = 50%				
Continuous variables				
Age, years	-0.018	0.129 ± 0.002	0.102 ± 0.001 ^a	0.094 ± 0.007 ^{a,}
SBP	0.492	0.115 ± 0.001	0.095 ± 0.001 ^a	$0.080 \pm 0.002 \ ^{\rm a}$
Charlson index	0.146	0.295 ± 0.001	0.239 ± 0.002 ^a	0.241 ± 0.009 ^a
TC/HDL-C ratio	3.139	0.279 ± 0.004	0.235 ± 0.003 ^a	0.183 ± 0.002 ^{a,d}
Hospital admission times	7.037	3.766 ± 0.12	3.199 ± 0.057 ^a	3.004 ± 0.246 ^{a,c}
Categorical variables				
Smoking	7.45%	0.335 ± 0.006	0.277 ± 0.015 ^a	0.267 ± 0.012 ^{a,c}
Hypertensive drugs	8.10%	0.368 ± 0.014	0.305 \pm 0.004 $^{\rm a}$	0.276 ± 0.005 ^{a,c}
Lipid Lowering drugs	9.99%	0.441 ± 0.015	0.319 ± 0.006 ^a	0.304 ± 0.009 ^{a,4}
Overweight	37.89%	1.135 ± 0.018	1.029 ± 0.019 ^a	0.850 ± 0.020 ^{a,c}
Sex	41.21%	1.149 ± 0.02	1.050 ± 0.013 ^a	1.007 ± 0.055 ^a

Notes

SBP Systolic Blood Pressure, TC Total Cholesterol, HDL-C High-Density Lipoprotein Cholesterol

Since NRMSE and PFC both followed normal distribution (Shapiro-Wilk normality test p value > 0.05), imputation errors of different methods were compared using one-way ANOVA. If p < 0.05, paired methods were compared using independent sample t-test;

^aThe mean imputation error is significantly lower than that of MICE (p < 0.05)

^bThe mean imputation error is significantly lower than that of GAIN (p < 0.05)

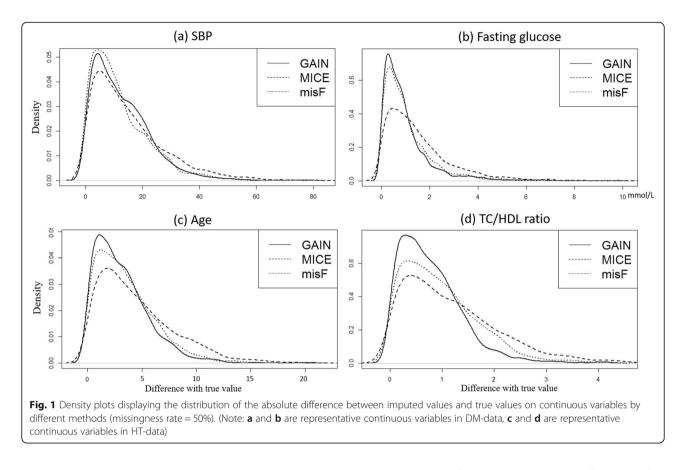
^cThe mean imputation error is significantly lower than that of missForest (p < 0.05)

subjects) and GAIN was the fastest for the larger sample (50,000 subjects). MissForest showed much longer computation times for all sample sizes compared to the other two methods. The computation time of missForest increased exponentially with increasing sample size.

Discussion

Missing data are inevitable in medical research and it is important that appropriate methods are used to solve this problem in order to make full use of the data and get unbiased inference. This study has introduced a novel imputation method, GAIN, and demonstrated its imputation accuracy and efficiency outperformed two commonly used methods (MICE and missForest). The major strength of this study was the use of two large real-world clinical datasets with mixed-type variables. To the best of our knowledge, this was the first study to evaluate the application of GAIN for the imputation of missing clinical data with mixed type variables.

Overall, GAIN showed similar imputation accuracy as missForest when the missingness rate was relatively low (20%) but performed better than missForest when the missingness rate was higher (50%). GAIN also had better accuracy for imputing skewed continuous variables and imbalanced categorical variables. Furthermore, the imputation time of GAIN increased only slightly with increasing sample size, making it the most efficient method for performing big data analytics on a sample size of more than 30,000.



Imputation performance and data characteristic

These findings matched those observed in an earlier study where GAIN outperformed other imputation methods on a cancer dataset in which all variables were continuous [20]. It is important to recognise that the imputations of mixed-type variables are challenging but essential for clinical research [10]. Our results provide preliminary evidence that GAIN is a suitable method for the imputation of missing clinical data with mixed type of variables, particularly those with highly skewed and imbalanced data.

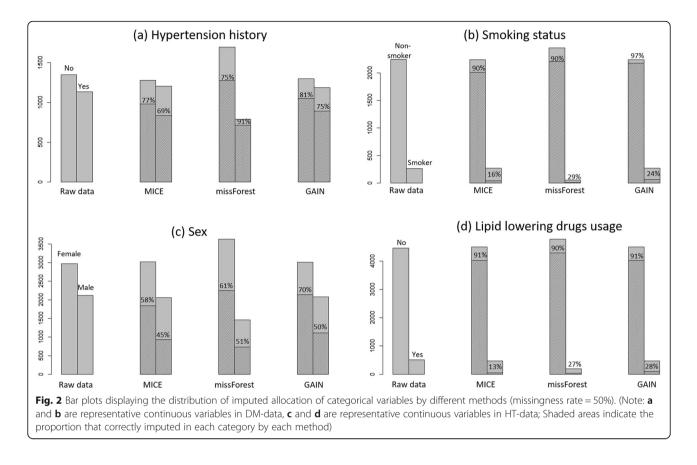
The results of this study also showed that, despite MICE being commonly used, there is still room for its improvement [14, 15]. As can be seen from the density plots, the default setting of MICE (pmm) replicated some observed extreme values to seek for the same distribution as the observed data, however, these extreme values might be far from the real values and lead to inaccuracy. On the other hand, missForest and GAIN, through machine learning, are more "moderate" and produced credible values, which are closer to the mean level of the observed data, yielding more accurate imputation results.

Imputation performance and missingness rate

It is recognized that data with a higher missing proportion are likely to increase further inference bias. There is no consensus on the maximum missing data rate that would allow for substitution by imputation since it is determined by various factors, including the missingness assumption, participation of auxiliary variables, data quality and also imputation methods [25]. In medical research and clinical trials, the rule of thumb for an acceptable missingness rate is 20% or less [26, 27], but much higher rates are commonly observed in real practice. For example, as shown in the two large real-world clinical datasets in this study, the data missingness rates of some variables were nearly 50%. In order to explore how the imputation accuracy would be affected by the data missingness rate, we evaluated the three methods on simulated data with missingness rates of 20 and 50%. It was found that GAIN was more resistant to the effects of a higher missingness rate. This is because the imputation power of GAIN depend not only on observed values but also on the feedback from the discriminator. GAIN therefore has the potential to accept a higher threshold of data missingness rate and maximize the use of research data.

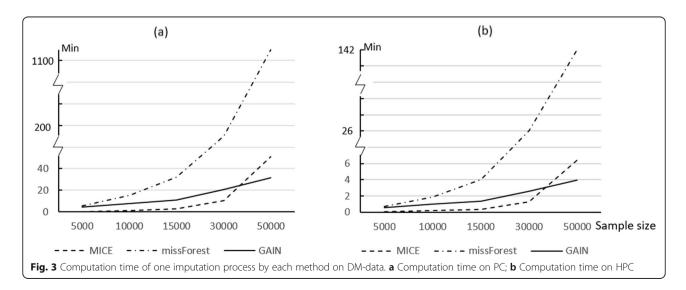
Computation time

In addition to the measures on accuracy, this study also recorded the computation time as a performance indicator. Computation time cannot be neglected, especially with the



large datasets in many cohort studies. GAIN stands out in its efficiency by virtue of its unique mechanism in which the number of parameters is relatively independent of the sample size.

Multiple imputation (MI) is recommended to avoid the uncertainty of single imputation. However, it will increase the computation time. In general, if MI is adopted, the imputation times (m) is at least 5 with some researchers using 10 or more [8]. MissForest will take approximately 8 days (200 h) of PC computation time to impute the missing data with a sample size of 50,000 with multiple imputation of 10 times. The computation time will also lengthen exponentially as the sample size increases. The utilization of HPC and parallel processing may save some time but may not be feasible in many settings.



Further implication for practice

There is no one best procedure to solve the problem of missing data in medical research. Indeed, the selected method will depend on the missingness assumption as well as auxiliary variables that could explain why the data is missing [28]. For example, complete case analysis might be preferable over MI in some situations [20]. Based on our findings, we would suggest taking into consideration missingness rate, variable distribution, and the expected computation time when choosing the appropriate imputation method. In addition, the use of more than one imputation method and sensitivity analysis could improve the reliability of the results.

Limitation

This study had a number of limitations. First and foremost, this study had only focused on the imputation accuracy but not post-imputation statistical inference effectiveness of different imputation methods. The goal of missing data imputation is to obtain statistically valid inferences from incomplete data rather than to re-create the true data. Van Buuren has pointed out that imputation is not prediction, and the method that best recovers the true data might be nonsensical or contain severe flaws [8]. Further studies should be conducted to evaluate these imputation methods with respect to postimputation statistical inferences. Second, a missingness rate of more than 50% was not simulated in this study as some researchers have suggested that a missingness rate of more than 50% is not acceptable for clinical studies [25]. Third, the variables included in this study were cross-sectional data, hence the results may not be generalizable to missing data problem in longitudinal studies with repeated observations.

Conclusion

Overall, when compared to MICE and missForest, GAIN showed better accuracy in the imputation of missing data in large real world clinical datasets, particularly for imbalanced and skewed data, and when the missingness rate was high (50%). GAIN also has outstanding computation speed in handling large samples (greater than 30, 000 subjects) and holds potential as an accurate and efficient method for missing data imputation in future big data clinical research.

Abbreviations

DM: Diabetes mellitus; GAIN: Generative adversarial imputation nets; GAN: Generative adversarial network; HDL: High-density lipoprotein; HPC: High performance computing; HT: Hypertension; LR: Logistic regression; MAR: Missing at random; MCAR: Missing completely at random; MICE: Multiple imputation by chained equations; MNAR: Missing not at random; NRMSE: Normalized root mean square error; PC: Personal computer; PFC: Proportion of falsely classified; pmm: Predictive mean matching; SBP: Systolic blood pressure; TC: Total cholesterol

Supplementary Information

The online version contains supplementary material available at https://doi. org/10.1186/s12874-021-01272-3.

Additional file 1: Supplementary Table 1. Characteristics Description of DM-data (N = 50,000). Supplementary Table 2. Characteristics Description of HT-data (N = 10,000). Supplementary Table 3. Computing devices and model hyper-parameters. Supplementary Figure 1. Performance of MICE and missForest at different iteration numbers.

Acknowledgements

We thank the Hong Kong Hospital Authority (HA) for the extraction of data from the HA computerized medical system. We would also like to show our gratitude to Prof. Liu Jia in Chinese Academy of Science, who inspired us in the conception of this study, and Dr. Ivy Mak, who helped to polish the English writing. The computations were performed using research computing facilities offered by Information Technology Services, the University of Hong Kong.

Authors' contributions

CLKL was the PI of the project and initialized this study. WD performed the data analysis and drafted the manuscript. DYTF, JY and EYFW oversaw the study design and data analysis. LEB, EHMT and CLKL participated in the results interpretation and manuscript revision. All the authors approved the final manuscript.

Funding

This work was supported by Food and Health Bureau of Hong Kong (Health and Medical Research Fund, 14151181)

Availability of data and materials

The data that support the findings of this study are available from Hong Kong Hospital Authority (HKHA) but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available. Data are however available from the authors upon reasonable request and with permission of HKHA.

Declarations

Ethics approval and consent to participate

The study was approved by the Institutional Review Board of the University of Hong Kong—the Hospital Authority Hong Kong West Cluster (reference number: UW 15–258).

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

Author details

¹Department of Family Medicine and Primary Care, Faculty of Medicine, University of Hong Kong, Hong Kong, Hong Kong SAR, China. ²School of Nursing, Faculty of Medicine, University of Hong Kong, Hong Kong, Hong Kong SAR, China. ³Electrical and Computer Engineering Department, University of California, Los Angeles, CA, USA. ⁴Department of Pharmacology and Pharmacy, Faculty of Medicine, University of Hong Kong, Hong Kong, Hong Kong SAR, China.

Received: 23 December 2020 Accepted: 6 April 2021 Published online: 20 April 2021

References

- Li P, Stuart EA, Allison DB. Multiple imputation: a flexible tool for handling missing DataMultiple ImputationMultiple imputation. JAMA. 2015;314(18): 1966–7. https://doi.org/10.1001/jama.2015.15281.
- Yoon J, Davtyan C, van der Schaar M. Discovery and clinical decision support for personalized healthcare. IEEE J Biomed Health Inform. 2017; 21(4):1133–45. https://doi.org/10.1109/JBHI.2016.2574857.

- Altman DG, Bland JM. Missing data. BMJ (Clinical research ed). 2007; 334(7590):424.
- Robinson KA, Dennison CR, Wayman DM, Pronovost PJ, Needham DM. Systematic review identifies number of strategies important for retaining study participants. J Clin Epidemiol. 2007;60(8):757.e1–e19.
- Hayati Rezvan P, Lee KJ, Simpson JA. The rise of multiple imputation: a review of the reporting and implementation of the method in medical research. BMC Med Res Methodol. 2015;15(1):30. https://doi.org/10.1186/s12 874-015-0022-1.
- Little RJA. In: Rubin DB, editor. Statistical analysis with missing data [electronic resource]. 2nd ed. Hoboken: Wiley; 2002. https://doi.org/10.1002/ 9781119013563.
- Graham JW. Missing data analysis: making it work in the real world. Annu Rev Psychol. 2009;60(1):549–76. https://doi.org/10.1146/annurev.psych.58.11 0405.085530.
- 8. Sv B. Flexible imputation of missing data. Boca Raton: CRC Press; 2012.
- Little R, Rubin D. Statistical analysis with missing data. 3rd ed. Hoboken: Wiley; 2019. https://doi.org/10.1002/9781119482260.
- Sterne JA, White IR, Carlin JB, Spratt M, Royston P, Kenward MG, et al. Multiple imputation for missing data in epidemiological and clinical research: potential and pitfalls. BMJ (Clinical research ed). 2009;338: b2393.
- Bhaskaran K, Smeeth L. What is the difference between missing completely at random and missing at random? Int J Epidemiol. 2014;43(4):1336–9. https://doi.org/10.1093/ije/dyu080.
- Pedersen AB, Mikkelsen EM, Cronin-Fenton D, Kristensen NR, Pham TM, Pedersen L, et al. Missing data and multiple imputation in clinical epidemiological research. Clin Epidemiol. 2017;9:157–66. https://doi.org/1 0.2147/CLEP.S129785.
- 13. Anon A. Multiple imputation by chained equations: what is it and how does it work? Int J Methods Psychiatr Res. 2011;20(1):40–9.
- Seaman SR, Bartlett JW, White IR. Multiple imputation of missing covariates with non-linear effects and interactions: an evaluation of statistical methods. BMC Med Res Methodol. 2012;12(1):46.
- 15. Tin KH. The random subspace method for constructing decision forests. IEEE Trans Pattern Anal Mach Intell. 1998;20(8):832–44.
- Stekhoven DJ, Bühlmann P. MissForest—non-parametric missing value imputation for mixed-type data. Bioinformatics. 2012;28(1):112–8. https://doi. org/10.1093/bioinformatics/btr597.
- Shah AD, Bartlett JW, Carpenter J, Nicholas O, Hemingway H. Comparison of random Forest and parametric imputation models for imputing missing data using MICE: a CALIBER study. Am J Epidemiol. 2014;179(6):764–74. https://doi.org/10.1093/aje/kwt312.
- Leon S, Benjamin W, Tim L. RenderGAN: generating realistic labeled data. Front Robot and Al. 2018;5.
- Bao J, Chen D, Wen F, Li H, Hua G. CVAE-GAN: Fine-Grained Image Generation through Asymmetric Training; 2017.
- Yoon J, Jordon J, van der Schaar M. GAIN: Missing Data Imputation using Generative Adversarial Nets; 2018.
- Wan EYF, Yu EYT, Chin WY, Fung CSC, Kwok RLP, Chao DVK, et al. Ten-year risk prediction models of complications and mortality of Chinese patients with diabetes mellitus in primary care in Hong Kong: a study protocol. BMJ Open. 2018;8(10):e023070.
- Fai Wan EY, Tak Yu EY, Chin WY, Fong DYT, Choi EPH, Tang EHM, et al. Burden of CKD and cardiovascular disease on life expectancy and health service utilization: a cohort study of Hong Kong Chinese hypertensive patients. J Am Soc Nephrol. 2019;30(10):1991–9. https://doi.org/10.1681/A SN.2018101037.
- Burgette LF, Reiter JP. Multiple imputation for missing data via sequential regression trees. Am J Epidemiol. 2010;172(9):1070–6. https://doi.org/10.1 093/aje/kwq260.
- 24. van Buuren S, Groothuis-Oudshoorn CGM. mice: Multivariate Imputation by Chained Equations in R. J Stat Softw. 2011;45(3):1548–7660.
- Royston P, Altman Douglas G, Marshall A, Holder RL. Comparison of techniques for handling missing covariate data within prognostic modelling studies: a simulation study. BMC Med Res Methodol. 2010;10(1):7.
- Schulz KF, Grimes DA. Sample size slippages in randomised trials: exclusions and the lost and wayward. Lancet. 2002;359(9308):781–5. https://doi.org/1 0.1016/S0140-6736(02)07882-0.
- Yeatts DS, Martin HR. What is missing from my missing data plan? Stroke. 2015;46(6):e130–e2. https://doi.org/10.1161/STROKEAHA.115.007984.

 Hughes RA, Heron J, Sterne JAC, Tilling K. Accounting for missing data in statistical analyses: multiple imputation is not always the answer. Int J Epidemiol. 2019;48(4):1294–304. https://doi.org/10.1093/ ije/dyz032.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Ready to submit your research? Choose BMC and benefit from:

- · fast, convenient online submission
- · thorough peer review by experienced researchers in your field
- rapid publication on acceptance
- support for research data, including large and complex data types
- gold Open Access which fosters wider collaboration and increased citations
- maximum visibility for your research: over 100M website views per year

At BMC, research is always in progress.

Learn more biomedcentral.com/submissions

